

APPLICATIONS TO ESTIMATION & FILTERING

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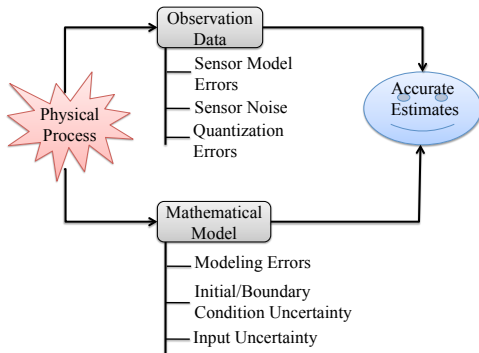
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MODEL-DATA FUSION



- The fusion of observational data with numerical simulation promises to provide **greater understanding of physical phenomenon** than either approach alone can achieve.
- The most critical challenge here is to provide a **quantitative assessment** of how closely our estimates reflect reality in the presence of model uncertainty as well as measurement errors and uncertainty.

DISCRETE DYNAMICS , DISCRETE MEASUREMENTS

$$\mathbf{x}_{k+1} = \mathbf{f}(\mathbf{x}_k) + \mathbf{v}_k$$

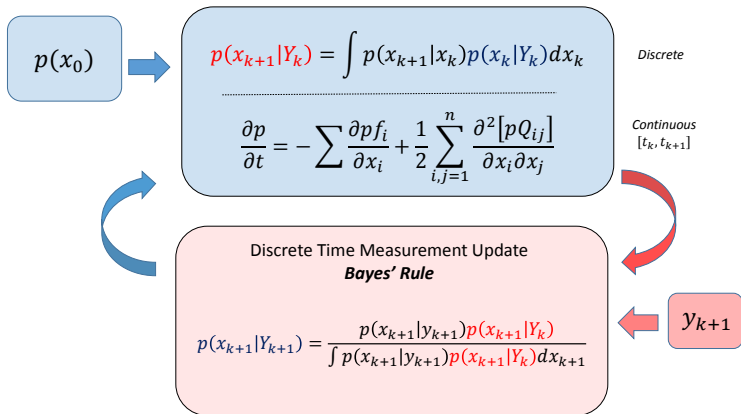
$$\mathbf{y}_{k+1} = \mathbf{h}(\mathbf{x}_{k+1}) + \boldsymbol{\omega}_{k+1}$$

CONTINUOUS DYNAMICS , DISCRETE MEASUREMENTS

$$d\mathbf{x} = \mathbf{f}(\mathbf{x}_t, t)dt + d\boldsymbol{\beta}_t \quad [t_k, t_{k+1}]$$

$$\mathbf{y}_{k+1} = \mathbf{h}(\mathbf{x}_{k+1}) + \boldsymbol{\omega}_{k+1}$$

IDEAL FILTER



LINEAR SYSTEM WITH GAUSSIAN UNCERTAINTY

DISCRETE TIME KALMAN FILTER

ASSUMPTIONS

- Initial conditions, Process noise and Measurement noise are **Gaussian** distributed and mutually independent processes.
- State and Measurement model equations are **linear**.

PROCESS AND MEASUREMENT MODEL

$$\mathbf{x}_{k+1} = F_k \mathbf{x}_k + \mathbf{v}_k$$

$$\mathbf{y}_{k+1} = H_{k+1} \mathbf{x}_{k+1} + \boldsymbol{\omega}_{k+1}$$

where

- $\mathbf{x}_0 \sim \mathcal{N}(\mathbf{x}_0 : \hat{\mathbf{x}}_0, P_0)$
- $\mathbf{v}_k \sim \mathcal{N}(\mathbf{v}_k : \mathbf{0}, Q_k)$
- $\boldsymbol{\omega}_{k+1} \sim \mathcal{N}(\boldsymbol{\omega}_{k+1} : \mathbf{0}, R_{k+1})$

TIME EVOLUTION

Chapman-Kolmogorov Equation: Given the state pdf at time step k

$$p(\mathbf{x}_k|Y_k) = \mathcal{N}(\mathbf{x}_k : \hat{\mathbf{x}}_{k|k}, P_{k|k})$$

where Y_k is set of measurement upto time k ,

$$p(\mathbf{x}_{k+1}|Y_k) = \int p(\mathbf{x}_{k+1}|\mathbf{x}_k) \cdot p(\mathbf{x}_k|Y_k) d\mathbf{x}_k$$

The state transition probability density $p(\mathbf{x}_{k+1}|\mathbf{x}_k)$ is given as

$$p(\mathbf{x}_{k+1}|\mathbf{x}_k) = \mathcal{N}(\mathbf{x}_{k+1} : F_k\mathbf{x}_k, Q_k)$$

$$p(\mathbf{x}_{k+1}|Y_k) = \int \mathcal{N}(\mathbf{x}_{k+1} : F_k\mathbf{x}_k, Q_k) \cdot \mathcal{N}(\mathbf{x}_k : \hat{\mathbf{x}}_{k|k}, P_{k|k}) d\mathbf{x}_k$$

TIME EVOLUTION

$$p(\mathbf{x}_{k+1}|Y_k) = \frac{1}{\sqrt{(2\pi)^n |P_{k+1|k}|}} \cdot \exp \left[-\frac{1}{2} (\mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1|k})^T P_{k+1|k}^{-1} (\mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1|k}) \right]$$

The prior pdf $p(\mathbf{x}_{k+1}|Y_k)$ has mean $\hat{\mathbf{x}}_{k+1|k}$ and covariance $P_{k+1|k}$ given as

$$\hat{\mathbf{x}}_{k+1|k} = F_k \hat{\mathbf{x}}_{k|k}$$

$$P_{k+1|k} = F_k P_{k|k} F_k^T + Q_k$$

Hence, the prior pdf remains Gaussian

As we know the pdf remains Gaussian, an easier approach is to directly compute the mean and covariance from the linear model equations as:

$$\hat{\mathbf{x}}_{k+1|k} = E[\mathbf{x}_{k+1}] = E[F_k x_k] + E[\mathbf{v}_k] = F_k E[x_k] = F_k \hat{\mathbf{x}}_{k|k}$$

$$\begin{aligned} P_{k+1|k} &= E[(\mathbf{x}_{k+1} - F_k \hat{\mathbf{x}}_{k|k})(\mathbf{x}_{k+1} - F_k \hat{\mathbf{x}}_{k|k})^T] = E[(F_k x_k + \mathbf{v}_k - F_k \hat{\mathbf{x}}_{k|k})(F_k x_k + \mathbf{v}_k - F_k \hat{\mathbf{x}}_{k|k})^T] \\ &= E[F_k (x_k - \hat{\mathbf{x}}_{k|k})(x_k - \hat{\mathbf{x}}_{k|k})^T F_k^T + \mathbf{v}_k \mathbf{v}_k^T] = F_k P_{k|k} F_k^T + Q_k \end{aligned}$$

MEASUREMENT UPDATE

Starting with the prior pdf at time step $k + 1$,

$$p(\mathbf{x}_{k+1}|Y_k) = \mathcal{N}(\mathbf{x}_{k+1} : \hat{\mathbf{x}}_{k+1|k}, P_{k+1|k})$$

the posterior pdf from Bayes' rule at the same time step is given by

$$p(\mathbf{x}_{k+1}|Y_{k+1}) = \frac{p(\mathbf{y}_{k+1}|\mathbf{x}_{k+1})p(\mathbf{x}_{k+1}|Y_k)}{\int p(\mathbf{y}_{k+1}|\mathbf{x}_{k+1})p(\mathbf{x}_{k+1}|Y_k) d\mathbf{x}_{k+1}}$$

where the measurement likelihood pdf $p(\mathbf{y}_{k+1}|\mathbf{x}_{k+1})$ can be derived from the measurement model equations as:

$$p(\mathbf{y}_{k+1}|\mathbf{x}_{k+1}) = \mathcal{N}(\mathbf{y}_{k+1} : H_{k+1}\mathbf{x}_{k+1}, R_{k+1})$$

Numerator of Bayes' rule is simplified as:

$$\begin{aligned}
 p(\mathbf{y}_{k+1}|\mathbf{x}_{k+1})p(\mathbf{x}_{k+1}|Y_k) &= \mathcal{N}(\mathbf{y}_{k+1} : H_{k+1}\mathbf{x}_{k+1}, R_{k+1})\mathcal{N}(\mathbf{x}_{k+1} : \hat{\mathbf{x}}_{k+1|k}, P_{k+1|k}) \\
 &= \frac{1}{\sqrt{|2\pi R_{k+1}|}} \cdot \frac{1}{\sqrt{|2\pi P_{k+1|k}|}} \cdot \exp \left[-\frac{1}{2} (\mathbf{y}_{k+1} - H_{k+1}\mathbf{x}_{k+1})^T R_{k+1}^{-1} (\mathbf{y}_{k+1} - H_{k+1}\mathbf{x}_{k+1}) \right. \\
 &\quad \left. - \frac{1}{2} (\mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1|k})^T P_{k+1|k}^{-1} (\mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1|k}) \right]
 \end{aligned}$$

The exponent is:

$$\begin{aligned}
 &\Rightarrow -\frac{1}{2} (\mathbf{y}_{k+1} - H_{k+1}\mathbf{x}_{k+1})^T R_{k+1}^{-1} (\mathbf{y}_{k+1} - H_{k+1}\mathbf{x}_{k+1}) - \frac{1}{2} (\mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1|k})^T P_{k+1|k}^{-1} (\mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1|k}) \\
 &\Rightarrow -\frac{1}{2} \left[\mathbf{y}_{k+1}^T R_{k+1}^{-1} \mathbf{y}_{k+1} + \mathbf{x}_{k+1}^T \underbrace{(P_{k+1|k}^{-1} + H_{k+1}^T R_{k+1}^{-1} H_{k+1})}_{A} \mathbf{x}_{k+1} \right. \\
 &\quad \left. - 2 \underbrace{(\mathbf{y}_{k+1}^T R_{k+1}^{-1} H_{k+1} + \hat{\mathbf{x}}_{k+1|k}^T P_{k+1|k}^{-1})}_{b^T} \mathbf{x}_{k+1} + \hat{\mathbf{x}}_{k+1|k}^T P_{k+1|k}^{-1} \hat{\mathbf{x}}_{k+1|k} \right] \\
 &\Rightarrow -\frac{1}{2} \left[\mathbf{y}_{k+1}^T R_{k+1}^{-1} \mathbf{y}_{k+1} + \mathbf{x}_{k+1}^T A \mathbf{x}_{k+1} - 2b^T \mathbf{x}_{k+1} + \hat{\mathbf{x}}_{k+1|k}^T P_{k+1|k}^{-1} \hat{\mathbf{x}}_{k+1|k} \right]
 \end{aligned}$$

LINEAR SYSTEM WITH GAUSSIAN UNCERTAINTY

DISCRETE TIME KALMAN FILTER

The denominator of Bayes' rule is given as:

$$\begin{aligned} \int p(\mathbf{y}_{k+1}|\mathbf{x}_{k+1})p(\mathbf{x}_{k+1}|Y_k) d\mathbf{x}_{k+1} &= \frac{1}{\sqrt{|2\pi R_{k+1}|}} \cdot \frac{1}{\sqrt{|2\pi P_{k+1|k}|}} \cdot \\ &\int \exp \left[-\frac{1}{2} \mathbf{y}_{k+1}^T R_{k+1}^{-1} \mathbf{y}_{k+1} - \frac{1}{2} \mathbf{x}_{k+1}^T A \mathbf{x}_{k+1} + b^T \mathbf{x}_{k+1} - \frac{1}{2} \hat{\mathbf{x}}_{k+1|k}^T P_{k+1|k}^{-1} \hat{\mathbf{x}}_{k+1|k} \right] d\mathbf{x}_{k+1} \\ &= \frac{1}{\sqrt{|2\pi R_{k+1}|}} \cdot \frac{1}{\sqrt{|2\pi P_{k+1|k}|}} \cdot \exp \left[-\frac{1}{2} \mathbf{y}_{k+1}^T R_{k+1}^{-1} \mathbf{y}_{k+1} - \frac{1}{2} \hat{\mathbf{x}}_{k+1|k}^T P_{k+1|k}^{-1} \hat{\mathbf{x}}_{k+1|k} \right] \\ &\quad \cdot \int \exp \left[-\frac{1}{2} \mathbf{x}_{k+1}^T A \mathbf{x}_{k+1} + b^T \mathbf{x}_{k+1} \right] d\mathbf{x}_{k+1} \\ &= \frac{\sqrt{|2\pi A^{-1}|}}{\sqrt{|2\pi R_{k+1}| |2\pi P_{k+1|k}|}} \cdot \exp \left[-\frac{1}{2} \mathbf{y}_{k+1}^T R_{k+1}^{-1} \mathbf{y}_{k+1} - \frac{1}{2} \hat{\mathbf{x}}_{k+1|k}^T P_{k+1|k}^{-1} \hat{\mathbf{x}}_{k+1|k} + \frac{1}{2} b^T A^{-T} b \right] \end{aligned}$$

The posterior pdf from Bayes' rule is given as:

$$\begin{aligned} p(\mathbf{x}_{k+1}|Y_k) &= \frac{1}{\sqrt{|2\pi A^{-1}|}} \cdot \exp \left[-\frac{1}{2} \mathbf{x}_{k+1}^T A \mathbf{x}_{k+1} + b^T \mathbf{x}_{k+1} - \frac{1}{2} b^T A^{-T} b \right] \\ &= \frac{1}{\sqrt{|2\pi A^{-1}|}} \cdot \exp \left[-\frac{1}{2} (\mathbf{x}_{k+1} - A^{-1}b)^T A (\mathbf{x}_{k+1} - A^{-1}b) \right] \end{aligned}$$

$$p(\mathbf{x}_{k+1}|Y_{k+1}) = \frac{1}{\sqrt{|2\pi A^{-1}|}} \cdot \exp \left[-\frac{1}{2} (\mathbf{x}_{k+1} - A^{-1}b)^T A (\mathbf{x}_{k+1} - A^{-1}b) \right]$$

$$b^T = (\mathbf{y}_{k+1}^T R_{k+1}^{-1} H_{k+1} + \hat{\mathbf{x}}_{k+1|k}^T P_{k+1|k}^{-1})$$

$$A = P_{k+1|k}^{-1} + H_{k+1}^T R_{k+1}^{-1} H_{k+1}$$

$$A^{-1} = P_{k+1|k} - \underbrace{P_{k+1|k} H_{k+1}^T (R_{k+1} + H_{k+1} P_{k+1|k} H_{k+1}^T)^{-1} H_{k+1} P_{k+1|k}}_{K_{k+1}}$$

$$= P_{k+1|k} - K_{k+1} H_{k+1} P_{k+1|k}$$

$$A^{-1}b = (P_{k+1|k} - K_{k+1} H_{k+1} P_{k+1|k}) (H_{k+1}^T R_{k+1}^{-1} \mathbf{y}_{k+1} + P_{k+1|k}^{-1} \hat{\mathbf{x}}_{k+1|k})$$

$$= \hat{\mathbf{x}}_{k+1|k} + K_{k+1} (\mathbf{y}_{k+1} - H_{k+1} \hat{\mathbf{x}}_{k+1|k})$$

Hence the posterior pdf still remains Gaussian even after Bayes' Rule update, with mean and covariance as

$$\hat{\mathbf{x}}_{k+1|k+1} = A^{-1}b = \hat{\mathbf{x}}_{k+1|k} + K_{k+1} (\mathbf{y}_{k+1} - H_{k+1} \hat{\mathbf{x}}_{k+1|k})$$

$$P_{k+1|k+1} = A^{-1} = P_{k+1|k} - K_{k+1} H_{k+1} P_{k+1|k}$$

In summary, starting from time step k with $p(\mathbf{x}_k|Y_k) = \mathcal{N}(\mathbf{x}_k : \hat{\mathbf{x}}_{k|k}, P_{k|k})$

TIME EVOLUTION

$$p(\mathbf{x}_{k+1}|Y_k) = \mathcal{N}(\mathbf{x}_{k+1} : \hat{\mathbf{x}}_{k+1|k}, P_{k+1|k})$$

$$\hat{\mathbf{x}}_{k+1|k} = F_k \hat{\mathbf{x}}_{k|k}$$

$$P_{k+1|k} = F_k P_{k|k} F^T + Q_k$$

MEASUREMENT UPDATE

$$p(\mathbf{x}_{k+1}|Y_{k+1}) = \mathcal{N}(\mathbf{x}_{k+1} : \hat{\mathbf{x}}_{k+1|k+1}, P_{k+1|k+1})$$

$$\hat{\mathbf{x}}_{k+1|k+1} = \hat{\mathbf{x}}_{k+1|k} + K_{k+1} (\mathbf{y}_{k+1} - H_{k+1} \hat{\mathbf{x}}_{k+1|k})$$

$$P_{k+1|k+1} = P_{k+1|k} - K_{k+1} H_{k+1} P_{k+1|k}$$

$$K_{k+1} = P_{k+1|k} H_{k+1}^T (R_{k+1} + H_{k+1} P_{k+1|k} H_{k+1}^T)^{-1}$$

The Kalman filter fuses the system dynamic model with measurement data in an optimal manner.

MINIMUM VARIANCE ESTIMATE

$$\min_{\hat{\mathbf{x}}_{k+1/k+1}} \text{Tr} \left\{ E [(\mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1/k+1}) (\mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1/k+1})^T] \right\}$$

i.e. find an estimate that minimizes the posterior variance

ESTIMATOR: KALMAN FILTER LIKE UPDATE

$$\hat{\mathbf{x}}_{k+1/k+1} = \hat{\mathbf{x}}_{k+1/k} + \mathbf{K}_{k+1} (\mathbf{y}_{k+1} - \hat{\mathbf{y}}_{k+1})$$

The assumed estimator is unbiased:

$$\begin{aligned} E[\hat{\mathbf{x}}_{k+1/k+1}] &= E[\hat{\mathbf{x}}_{k+1/k} + \mathbf{K}_{k+1} (\mathbf{y}_{k+1} - \hat{\mathbf{y}}_{k+1})] = E[\hat{\mathbf{x}}_{k+1/k}] + \mathbf{K}_{k+1} (E[\mathbf{y}_{k+1}] - E[\hat{\mathbf{y}}_{k+1}]) \\ &= F_k E[\hat{\mathbf{x}}_{k|k}] + \mathbf{K}_{k+1} (H_{k+1} E[\mathbf{x}_{k+1}] - H_{k+1} E[\hat{\mathbf{x}}_{k+1/k}]) \\ &= E[\mathbf{x}_{k+1}] + \mathbf{K}_{k+1} (H_{k+1} F \hat{\mathbf{x}}_{k|k} - H_{k+1} F \hat{\mathbf{x}}_{k|k}) = E[\mathbf{x}_{k+1}] \end{aligned}$$

LINEAR SYSTEM WITH GAUSSIAN UNCERTAINTY

DISCRETE TIME MINIMUM VARIANCE APPROACH FOR KALMAN FILTER

$$\min_{\mathbf{K}_{k+1}} Tr \left\{ E[(\mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1/k+1})(\mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1/k+1})^T] \right\}$$

with $\hat{\mathbf{x}}_{k+1/k+1} = \hat{\mathbf{x}}_{k+1/k} + \mathbf{K}_{k+1}(\mathbf{y}_{k+1} - \hat{\mathbf{y}}_{k+1})$

$$\min_{\mathbf{K}_{k+1}} Tr \left\{ E[(\mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1/k} - \mathbf{K}_{k+1}(\mathbf{y}_{k+1} - \hat{\mathbf{y}}_{k+1}))(\mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1/k} - \mathbf{K}_{k+1}(\mathbf{y}_{k+1} - \hat{\mathbf{y}}_{k+1}))^T] \right\}$$

$$\min_{\mathbf{K}_{k+1}} Tr \left\{ E[(\mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1/k})(\mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1/k})^T] + \mathbf{K}_{k+1} E[(\mathbf{y}_{k+1} - \hat{\mathbf{y}}_{k+1})(\mathbf{y}_{k+1} - \hat{\mathbf{y}}_{k+1})^T] \mathbf{K}_{k+1}^T \right. \\ \left. - \mathbf{K}_{k+1} E[(\mathbf{y}_{k+1} - \hat{\mathbf{y}}_{k+1})(\mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1/k})^T] - E[(\mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1/k})(\mathbf{y}_{k+1} - \hat{\mathbf{y}}_{k+1})^T] \mathbf{K}_{k+1}^T \right\}$$

$$\min_{\mathbf{K}_{k+1}} Tr \left\{ P_{k+1|k} + \mathbf{K}_{k+1} P_{k+1|k}^y \mathbf{K}_{k+1}^T - \mathbf{K}_{k+1} P_{k+1|k}^{yx} - P_{k+1|k}^{xy} \mathbf{K}_{k+1}^T \right\}$$

The optimal gain \mathbf{K}_{k+1} is given by

$$\mathbf{K}_{k+1} = \mathbf{P}_{k+1}^{xy} (\mathbf{P}_{k+1|k}^y)^{-1}$$

- no assumptions on the state pdf.
- All the expectations are with respect to the prior pdf $p(\mathbf{x}_{k+1}|Y_k)$.

LINEAR SYSTEM WITH GAUSSIAN UNCERTAINTY

DISCRETE TIME MINIMUM VARIANCE APPROACH FOR KALMAN FILTER

$$\begin{aligned}\hat{\mathbf{y}}_{k+1} &= E[\mathbf{y}_{k+1}] = E[H_{k+1}\mathbf{x}_{k+1} + \boldsymbol{\omega}_{k+1}] = H_{k+1}E[\mathbf{x}_{k+1}] = H_{k+1}\hat{\mathbf{x}}_{k+1|k} \\ P_{k+1|k}^y &= E[(\mathbf{y}_{k+1} - \hat{\mathbf{y}}_{k+1})(\mathbf{y}_{k+1} - \hat{\mathbf{y}}_{k+1})^T] \\ &= E[(H_{k+1}\mathbf{x}_{k+1} + \boldsymbol{\omega}_{k+1} - H_{k+1}\hat{\mathbf{x}}_{k+1|k})(H_{k+1}\mathbf{x}_{k+1} + \boldsymbol{\omega}_{k+1} - H_{k+1}\hat{\mathbf{x}}_{k+1|k})^T] \\ &= H_{k+1}E[(\mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1|k})(\mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1|k})^T]H_{k+1}^T + E[\boldsymbol{\omega}_{k+1}\boldsymbol{\omega}_{k+1}^T] \\ &= H_{k+1}P_{k+1|k}H_{k+1}^T + R_{k+1} \\ P_{k+1}^{xy} &= E[(\mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1|k})(\mathbf{y}_{k+1} - \hat{\mathbf{y}}_{k+1})^T] \\ &= E[(\mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1|k})(H_{k+1}(\mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1|k}) + \boldsymbol{\omega}_{k+1})^T] \\ &= E[(\mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1|k})(\mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1|k})^T]H_{k+1} = P_{k+1|k}H_{k+1}^T\end{aligned}$$

The Kalman Filter gain is then given as:

$$\mathbf{K}_{k+1} = \mathbf{P}_{k+1}^{xy}(\mathbf{P}_{k+1|k}^y)^{-1} = P_{k+1|k}H_{k+1}^T(H_{k+1}P_{k+1|k}H_{k+1}^T + R_{k+1})^{-1}$$

Minimum variance estimator is same as Kalman Filter and hence optimal for linear system with Gaussian pdfs

Extended Kalman Filter: Linearization of system model equations

PROCESS AND MEASUREMENT MODEL

$$\mathbf{x}_{k+1} = f_k(\mathbf{x}_k) + \mathbf{v}_k$$

$$\mathbf{y}_{k+1} = h_{k+1}(\mathbf{x}_{k+1}) + \boldsymbol{\omega}_{k+1}$$

Starting with the most recent estimates at time k i.e. $\hat{\mathbf{x}}_{k|k}$ and $P_{k|k}$, the posterior estimates are approximated as

$$\begin{aligned} E[\mathbf{x}_{k+1}] &= E[f_k(\mathbf{x}_k) + \mathbf{v}_k] = E[f_k(\mathbf{x}_k)] + E[\mathbf{v}_k] = E[f_k(\mathbf{x}_k)] \\ &= E[f_k(\mathbf{x}_k)] = E[f_k(\hat{\mathbf{x}}_{k|k}) + F_k(\mathbf{x}_k - \hat{\mathbf{x}}_{k|k}) + \dots] \approx f_k(\hat{\mathbf{x}}_{k|k}) \end{aligned}$$

using the approximation

$$f_k(\mathbf{x}_k) - f_k(\hat{\mathbf{x}}_{k|k}) \approx F_k(\mathbf{x}_k - \hat{\mathbf{x}}_{k|k}) \quad \text{where} \quad F_k \equiv \left. \frac{\partial f_k}{\partial \mathbf{x}_k} \right|_{\mathbf{x}_k = \hat{\mathbf{x}}_{k|k}}$$

$$\begin{aligned} & E[(\mathbf{x}_{k+1} - f_k(\hat{\mathbf{x}}_{k|k}))(\mathbf{x}_{k+1} - f_k(\hat{\mathbf{x}}_{k|k}))^T] \\ &= E[(f_k(\mathbf{x}_k) + \mathbf{v}_k - f_k(\hat{\mathbf{x}}_{k|k}))(f_k(\mathbf{x}_k) + \mathbf{v}_k - f_k(\hat{\mathbf{x}}_{k|k}))^T] \\ &= E\left[\left(F_k(\mathbf{x}_k - \hat{\mathbf{x}}_{k|k}) + \mathbf{v}_k\right)\left(F_k(\mathbf{x}_k - \hat{\mathbf{x}}_{k|k}) + \mathbf{v}_k\right)^T\right] \\ &= F_k E[(\mathbf{x}_k - \hat{\mathbf{x}}_{k|k})(\mathbf{x}_k - \hat{\mathbf{x}}_{k|k})^T] F_k^T + E[\mathbf{v}_k \mathbf{v}_k^T] = F_k P_{k|k} F_k^T + Q_k \end{aligned}$$

Only mean and covariance are propagated

EKF: TIME EVOLUTION STEP

$$\hat{\mathbf{x}}_{k+1|k} = f_k(\hat{\mathbf{x}}_{k|k})$$

$$P_{k+1|k} = F_k P_{k|k} F_k^T + Q_k$$

Minimum variance estimator with Kalman Filter like update for the EKF:

$$\min_{\mathbf{K}_{k+1}} Tr \left\{ E[(\mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1/k+1})(\mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1/k+1})^T] \right\}$$

with $\hat{\mathbf{x}}_{k+1/k+1} = \hat{\mathbf{x}}_{k+1/k} + \mathbf{K}_{k+1}(\mathbf{y}_{k+1} - \hat{\mathbf{y}}_{k+1})$

$$\min_{\mathbf{K}_{k+1}} Tr \left\{ P_{k+1|k} + \mathbf{K}_{k+1} P_{k+1|k}^y \mathbf{K}_{k+1}^T - \mathbf{K}_{k+1} P_{k+1|k}^{yx} - P_{k+1|k}^{xy} \mathbf{K}_{k+1}^T \right\}$$

The optimal gain \mathbf{K}_{k+1} is given by

$$\mathbf{K}_{k+1} = \mathbf{P}_{k+1}^{xy} (\mathbf{P}_{k+1/k}^y)^{-1}$$

Using Taylor Series Expansion of $h(\mathbf{x}_{k+1})$ about the current estimate at time $k+1$ i.e. $\hat{\mathbf{x}}_{k+1|k}$

$$h_{k+1}(\mathbf{x}_{k+1}) = h_{k+1}(\hat{\mathbf{x}}_{k+1|k}) + H_{k+1}(\mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1|k}) \quad \text{where} \quad H_{k+1} \equiv \frac{\partial h_{k+1}}{\partial \mathbf{x}_{k+1}} \Big|_{\mathbf{x}_{k+1} = \hat{\mathbf{x}}_{k+1|k}}$$

$$\hat{\mathbf{y}}_{k+1} = E[\mathbf{y}_{k+1}] = E[h_{k+1}(\mathbf{x}_{k+1})] + E[\boldsymbol{\omega}_{k+1}] \approx h_{k+1}(\hat{\mathbf{x}}_{k+1|k})$$

$$\begin{aligned}
 P_{k+1/k}^y &= E[(\mathbf{y}_{k+1} - \hat{\mathbf{y}}_{k+1})(\mathbf{y}_{k+1} - \hat{\mathbf{y}}_{k+1})^T] \\
 &\approx E[(h_{k+1}(\mathbf{x}_{k+1}) + \boldsymbol{\omega}_{k+1} - h_{k+1}(\hat{\mathbf{x}}_{k+1|k})) (h_{k+1}(\mathbf{x}_{k+1}) + \boldsymbol{\omega}_{k+1} - h_{k+1}(\hat{\mathbf{x}}_{k+1|k}))^T] \\
 &= H_{k+1} E[(\mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1|k})(\mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1|k})^T] H_{k+1}^T + E[\boldsymbol{\omega}_{k+1} \boldsymbol{\omega}_{k+1}^T] \\
 &= H_{k+1} P_{k+1|k} H_{k+1}^T + R_{k+1} \\
 P_{k+1}^{xy} &= E[(\mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1|k})(\mathbf{y}_{k+1} - \hat{\mathbf{y}}_{k+1})^T] \\
 &\approx E[(\mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1|k})(h_{k+1}(\mathbf{x}_{k+1}) + \boldsymbol{\omega}_{k+1} - h_{k+1}(\hat{\mathbf{x}}_{k+1|k}))^T] \\
 &= E[(\mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1|k})(\mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1|k})^T] H_{k+1} = P_{k+1|k} H_{k+1}^T
 \end{aligned}$$

The Extended Kalman Filter gain is then given as

$$\mathbf{K}_{k+1} = \mathbf{P}_{k+1}^{xy} (\mathbf{P}_{k+1}^y)^{-1} = P_{k+1|k} H_{k+1}^T (H_{k+1} P_{k+1|k} H_{k+1}^T + R_{k+1})^{-1}$$

NONLINEAR SYSTEM WITH GAUSSIAN UNCERTAINTY

DISCRETE TIME EXTENDED KALMAN FILTER

In summary,

TIME EVOLUTION

$$\begin{aligned}\hat{\mathbf{x}}_{k+1|k} &= f_k(\hat{\mathbf{x}}_{k|k}) \\ P_{k+1|k} &= F_k P_{k|k} F_k^T + Q_k \\ F_k &\equiv \left. \frac{\partial f_k}{\partial \mathbf{x}_k} \right|_{\mathbf{x}_k = \hat{\mathbf{x}}_{k|k}}\end{aligned}$$

MEASUREMENT UPDATE

$$\begin{aligned}\hat{\mathbf{x}}_{k+1/k+1} &= \hat{\mathbf{x}}_{k+1/k} + \mathbf{K}_{k+1}(\mathbf{y}_{k+1} - h_{k+1}(\hat{\mathbf{x}}_{k+1/k})) \\ P_{k+1|k+1} &= P_{k+1|k} - \mathbf{K}_{k+1} H_{k+1} P_{k+1|k} \\ \mathbf{K}_{k+1} &= P_{k+1|k} H_{k+1}^T (H_{k+1} P_{k+1|k} H_{k+1}^T + R_{k+1})^{-1} \\ H_{k+1} &\equiv \left. \frac{\partial h_{k+1}}{\partial \mathbf{x}_{k+1}} \right|_{\mathbf{x}_{k+1} = \hat{\mathbf{x}}_{k+1|k}}\end{aligned}$$

- Only mean and covariance are propagated and updated.
- All expectation expression $E[\cdot]$ evaluated by linearizations \Rightarrow analytical expressions
- Estimates can quickly diverge due to linearizations involved.

MOTIVATION FOR UNSCENTED AND QUADRATURE KALMAN FILTER

- \Rightarrow **Avoid linearization** altogether and evaluate these expectation integrals directly using **appropriate quadrature scheme**.
- *Unscented Transform, Gauss-hermite Quadratures or Conjugate Unscented Transform*

DISCRETE DYNAMIC SYSTEM

$$\begin{aligned}\mathbf{x}_{k+1} &= \mathbf{f}(\mathbf{x}_k) + \mathbf{v}_k \\ \mathbf{y}_{k+1} &= \mathbf{h}(\mathbf{x}_{k+1}) + \boldsymbol{\omega}_{k+1}\end{aligned}$$

Evolution of the exact conditional pdf is given by two stages

TIME EVOLUTION STEP

$$p(\mathbf{x}_{k+1}|\mathbf{Y}_k) = \int \mathcal{N}(\mathbf{x}_{k+1} : \mathbf{f}(\mathbf{x}_k), \mathbf{Q}_k) \cdot p(\mathbf{x}_k|\mathbf{Y}_k) d\mathbf{x}_k$$

MEASUREMENT UPDATE STEP

$$p(\mathbf{x}_{k+1}|\mathbf{Y}_{k+1}) = \frac{\mathcal{N}(\mathbf{y}_{k+1} : \mathbf{h}(\mathbf{x}_{k+1}), \mathbf{R}_{k+1}) \cdot p(\mathbf{x}_{k+1}|\mathbf{Y}_k)}{\int \mathcal{N}(\mathbf{y}_{k+1} : \mathbf{h}(\mathbf{x}_{k+1}), \mathbf{R}_{k+1}) \cdot p(\mathbf{x}_{k+1}|\mathbf{Y}_k) d\mathbf{x}_{k+1}}$$

Gaussian Approximated conditional pdfs:

TIME EVOLUTION STEP

$$p(\mathbf{x}_{k+1}|\mathbf{Y}_k) = \int \mathcal{N}(\mathbf{x}_{k+1} : \mathbf{f}(\mathbf{x}_k), \mathbf{Q}_k) \cdot p(\mathbf{x}_k|\mathbf{Y}_k) d\mathbf{x}_k$$

$$\text{Mean} : \int \mathbf{x}_{k+1} p(\mathbf{x}_{k+1}|\mathbf{Y}_k) d\mathbf{x}_{k+1} \Rightarrow \hat{\mathbf{x}}_{k+1/k}$$

$$\text{Covariance} : \int (\mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1/k})(\mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1/k})^T p(\mathbf{x}_{k+1}|\mathbf{Y}_k) d\mathbf{x}_{k+1} \Rightarrow \mathbf{P}_{k+1/k}$$

$$\hat{\mathbf{x}}_{k+1/k} = \int \mathbf{f}(\mathbf{x}_k) \mathcal{N}(\mathbf{x}_k : \hat{\mathbf{x}}_{k/k}, \mathbf{P}_{k/k} | \mathbf{Y}_k) d\mathbf{x}_k$$

$$\mathbf{P}_{k+1/k} = \int \mathbf{f}(\mathbf{x}_k) \mathbf{f}(\mathbf{x}_k)^T \mathcal{N}(\mathbf{x}_k : \hat{\mathbf{x}}_{k/k}, \mathbf{P}_{k/k} | \mathbf{Y}_k) d\mathbf{x}_k - \hat{\mathbf{x}}_{k+1/k} \hat{\mathbf{x}}_{k+1/k}^T + \mathbf{Q}_k$$

GAUSSIAN PDF AT $k+1$

$$p(\mathbf{x}_{k+1}|\mathbf{Y}_k) \approx \mathcal{N}(\mathbf{x}_{k+1} : \hat{\mathbf{x}}_{k+1/k}, \mathbf{P}_{k+1/k})$$

Gaussian Approximated conditional pdfs:

TIME EVOLUTION STEP

$$p(\mathbf{x}_{k+1}|\mathbf{Y}_k) = \int \mathcal{N}(\mathbf{x}_{k+1} : \mathbf{f}(\mathbf{x}_k), \mathbf{Q}_k) \cdot \mathcal{N}(\mathbf{x}_k : \hat{\mathbf{x}}_{k/k}, \mathbf{P}_{k/k}) d\mathbf{x}_k$$

$$\text{Mean} : \int \mathbf{x}_{k+1} p(\mathbf{x}_{k+1}|\mathbf{Y}_k) d\mathbf{x}_{k+1} \Rightarrow \hat{\mathbf{x}}_{k+1/k}$$

$$\text{Covariance} : \int (\mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1/k})(\mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1/k})^T p(\mathbf{x}_{k+1}|\mathbf{Y}_k) d\mathbf{x}_{k+1} \Rightarrow \mathbf{P}_{k+1/k}$$

$$\hat{\mathbf{x}}_{k+1/k} = \int \mathbf{f}(\mathbf{x}_k) \mathcal{N}(\mathbf{x}_k : \hat{\mathbf{x}}_{k/k}, \mathbf{P}_{k/k} | \mathbf{Y}_k) d\mathbf{x}_k$$

$$\mathbf{P}_{k+1/k} = \int \mathbf{f}(\mathbf{x}_k) \mathbf{f}(\mathbf{x}_k)^T \mathcal{N}(\mathbf{x}_k : \hat{\mathbf{x}}_{k/k}, \mathbf{P}_{k/k} | \mathbf{Y}_k) d\mathbf{x}_k - \hat{\mathbf{x}}_{k+1/k} \hat{\mathbf{x}}_{k+1/k}^T + \mathbf{Q}_k$$

GAUSSIAN PDF AT $k+1$

$$p(\mathbf{x}_{k+1}|\mathbf{Y}_k) \approx \mathcal{N}(\mathbf{x}_{k+1} : \hat{\mathbf{x}}_{k+1/k}, \mathbf{P}_{k+1/k})$$

NONLINEAR FILTERING- MEASUREMENT UPDATE

USING QUADRATURES/CUBATURES/SIGMA POINTS

Assuming a Kalman Filter like update

$$\hat{\mathbf{x}}_{k+1/k+1} = \hat{\mathbf{x}}_{k+1/k} + \mathbf{K}_{k+1}(\mathbf{y}_{k+1} - \hat{\mathbf{y}}_{k+1})$$

$$\min_{\mathbf{K}_{k+1}} Tr \left\{ E[(\mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1/k+1})(\mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1/k+1})^T] \right\} \quad (1)$$

$$= Tr \left\{ \mathbf{P}_{k+1/k} \right\} - Tr \left\{ \mathbf{P}_{k+1}^{xy} \mathbf{K}_{k+1}^T \right\} - Tr \left\{ \mathbf{K}_{k+1} (\mathbf{P}_{k+1}^{xy})^T \right\} + Tr \left\{ \mathbf{K}_{k+1} \mathbf{P}_{k+1/k}^y \mathbf{K}_{k+1}^T \right\} \quad (2)$$

The Kalman gain is given as

$$\mathbf{K}_{k+1} = \mathbf{P}_{k+1}^{xy} (\mathbf{P}_{k+1/k}^y)^{-1}$$

KALMAN FILTER LIKE UPDATE $p(\mathbf{x}_{k+1}|\mathbf{y}_{k+1})$

$$\hat{\mathbf{x}}_{k+1/k+1} = \hat{\mathbf{x}}_{k+1/k} + \mathbf{K}_{k+1}(\mathbf{y}_{k+1} - \hat{\mathbf{y}}_{k+1})$$

$$\mathbf{P}_{k+1/k+1} = \mathbf{P}_{k+1/k} - \mathbf{K}_{k+1} (\mathbf{P}_{k+1/k}^{xy})^T$$

$$\mathbf{K}_{k+1} = \mathbf{P}_{k+1}^{xy} (\mathbf{P}_{k+1/k}^y)^{-1}$$

$$p(\mathbf{x}_{k+1}|\mathbf{Y}_{k+1}) \approx \mathcal{N}(\mathbf{x}_{k+1} : \hat{\mathbf{x}}_{k+1/k+1}, \mathbf{P}_{k+1/k+1})$$

The integrals to be evaluated are summarised as

TIME EVOLUTION STEP

$$\hat{\mathbf{x}}_{k+1/k} = \int \mathbf{f}(\mathbf{x}_k) \mathcal{N}(\mathbf{x}_k : \hat{\mathbf{x}}_{k/k}, \mathbf{P}_{k/k} | \mathbf{Y}_k) d\mathbf{x}_k$$

$$\mathbf{P}_{k+1/k} = \int \mathbf{f}(\mathbf{x}_k) \mathbf{f}(\mathbf{x}_k)^T \mathcal{N}(\mathbf{x}_k : \hat{\mathbf{x}}_{k/k}, \mathbf{P}_{k/k} | \mathbf{Y}_k) d\mathbf{x}_k \\ - \hat{\mathbf{x}}_{k+1/k} \hat{\mathbf{x}}_{k+1/k}^T + \mathbf{Q}_k$$

MEASUREMENT UPDATE STEP

$$\hat{\mathbf{y}}_{k+1} = \int \mathbf{h}(\mathbf{x}_{k+1/k}) \mathcal{N}(\mathbf{x}_{k+1/k} : \hat{\mathbf{x}}_{k+1/k}, \mathbf{P}_{k+1/k}) d\mathbf{x}_{k+1/k}$$

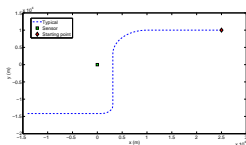
$$\mathbf{P}_{k+1/k}^y = \int \mathbf{h}(\mathbf{x}_{k+1/k}) \mathbf{h}(\mathbf{x}_{k+1/k})^T \mathcal{N}(\mathbf{x}_{k+1/k} : \hat{\mathbf{x}}_{k+1/k}, \mathbf{P}_{k+1/k}) d\mathbf{x}_{k+1/k} \\ - \hat{\mathbf{y}}_{k+1} \hat{\mathbf{y}}_{k+1}^T + \mathbf{R}_{k+1}$$

$$\mathbf{P}_{k+1}^{xy} = \int \mathbf{x}_{k+1/k} \mathbf{h}(\mathbf{x}_{k+1/k})^T \mathcal{N}(\mathbf{x}_{k+1/k} : \hat{\mathbf{x}}_{k+1/k}, \mathbf{P}_{k+1/k}) d\mathbf{x}_{k+1/k} \\ - \hat{\mathbf{x}}_{k+1/k} \hat{\mathbf{y}}_{k+1}^T$$

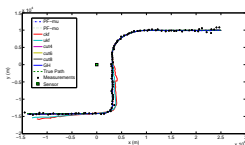
$(\mathbf{X}; \mathbf{w}) \in \left\{ \begin{array}{l} \textit{Monte Carlo samples} \\ \textit{Gauss-Hermite Product rule} \\ \textit{Sparse Grid Gauss-Hermite Quadrature rule} \\ \textit{Unscented Transform} \\ \textit{Cubature Kalman Filter} \\ \textit{Conjugate Unscented Transform} \\ \textit{Minimal Cubature rules} \end{array} \right.$

NONLINEAR FILTERING: NUMERICAL EXAMPLE

AIR TRAFFIC SCENARIO



(a) Typical Aeroplane Trajectory



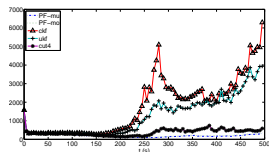
(b) Estimated Aeroplane Trajectory

Figure: Air Traffic Scenario

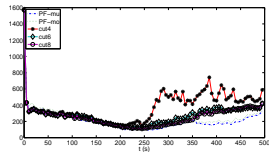
$$x_k = \begin{bmatrix} 1 & \frac{\sin(\Omega T)}{\Omega} & 0 & -\frac{1 - \cos(\Omega T)}{\Omega} \\ 0 & \cos(\Omega T) & 0 & -\frac{\sin(\Omega T)}{\Omega} \\ 0 & \frac{1 - \cos(\Omega T)}{\Omega} & 1 & \frac{\sin(\Omega T)}{\Omega} \\ 0 & \sin(\Omega T) & 0 & \cos(\Omega T) \\ 0 & 0 & 0 & 0 \end{bmatrix} x_{k-1} + v_{k-1}, \quad \begin{bmatrix} r_k \\ \theta_k \end{bmatrix} = \begin{bmatrix} \sqrt{(\xi_k)^2 + (\eta_k)^2} \\ \tan^{-1}(\frac{\eta_k}{\xi_k}) \end{bmatrix} + \omega_k$$

TABLE: Comparison of 2-norms of RMSE in position, velocity and turn rate for $T = 5s$ for Air Traffic Problem.

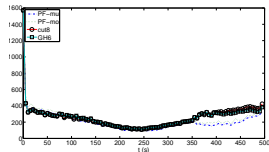
$\ RMSE\ _2$ in	PF - mean	CKF	UKF	CUT4	CUT6	CUT8
Position	115.47	989.15	685.90	245.30	138.82	135.89
Velocity	24.16	17330.23	12849.90	6127.86	2153.53	34.73
Ω	0.0393	2.873	2.396	1.329	0.636	0.090
No. of points	5000	10	11	42	83	355



(a) PF,CKF,UKF,CUT4



(b) PF,CUT4,CUT6,CUT8



(c) PF,CUT8,GH6

NONLINEAR FILTERING: NUMERICAL EXAMPLE

AIR TRAFFIC SCENARIO

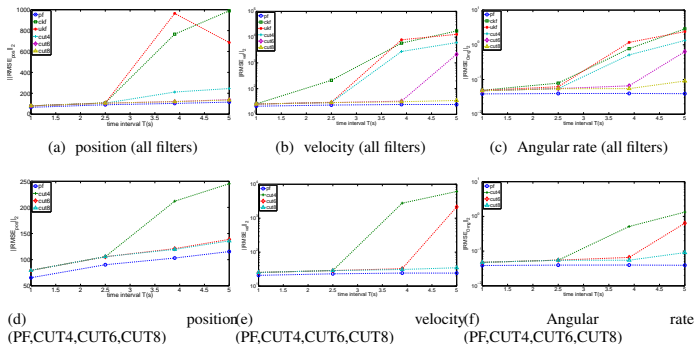
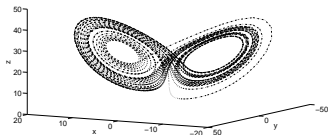


Figure: State Estimation Error vs. the Measurement Time Interval, T for the Air Traffic problem

NONLINEAR FILTERING: NUMERICAL EXAMPLE

LORENZ MODEL WITH PARAMETER UNCERTAINTY



$$\dot{x} = \sigma(y - x)$$

$$\dot{y} = \rho x - y - xz$$

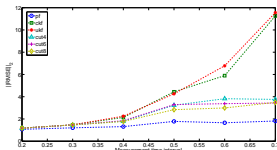
$$\dot{z} = xy - \beta z$$

$\sigma = 10, \rho = 28$ and $\beta = 8/3$. For uncertain σ and ρ , the appended state vector is $[x, y, z, \sigma, \rho]^T$ with mean and covariance

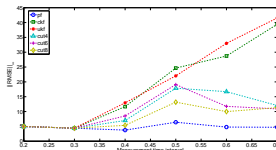
$$\mu_0 = [1.50887, -1.531271, 25.46091, 10, 28]^T;$$

$$P_0 = \text{Diag}([4, 4, 4, 2, 4]^T)$$

Figure: Lorenz system



(a) 2 norm of RMSE



(b) max of RMSE

Figure: Comparison of Filters for Lorenz model with varying measurement time intervals T

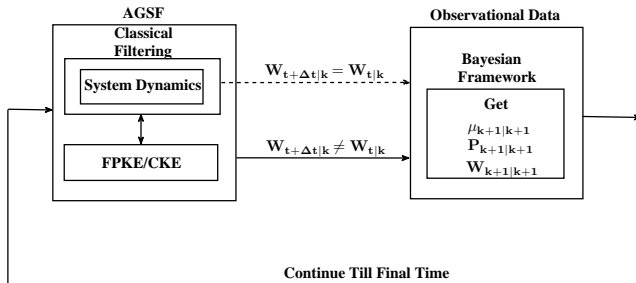
TABLE: Comparison of RMSE for various filters

	PF - mean	CKF	UKF	CUT4	CUT6	CUT8
$\ RMSE\ _2$ with $Q = 0.002I_{5 \times 5}$	14.7986	0.8945	0.8934	0.8974	0.9004	0.9002
$\ RMSE\ _2$ with $Q = 0.005I_{5 \times 5}$	1.0826	6.3419	4.2319	1.9162	1.8156	1.7225
No. of pts	5000	10	11	42	83	355

Measurement Model:

$$\mathbf{z}_k = \mathbf{h}(t_k, \mathbf{x}_k) + \mathbf{v}_k$$

$$\mathbf{v}_k \sim \mathcal{N}(\mathbf{0}, \mathbf{R}_k)$$



- Use EKF or UKF measurement equations to get $\mu_{k+1|k+1}$, $\mathbf{P}_{k+1|k+1}$.
- Use Bayes' rule to get the new weights.

Consider the following measurement model in discrete time

$$\mathbf{z}_k = \mathbf{h}(t_k, \mathbf{x}_k) + \mathbf{v}_k \quad \mathbf{v}_k \sim \mathcal{N}(\mathbf{0}, \mathbf{R}_k) \quad (3)$$

Then the Bayes' rule updates the weight according to the following equation:

$$\mathbf{w}_{k+1|k+1}^i = \frac{w_{k+1|k}^i \gamma^i}{\sum_{i=1}^N w_{k+1|k}^i \gamma^i} \quad i = 1, 2, \dots, N \quad (4)$$

where N = total number of Gaussian components and

$\gamma^i \sim \mathcal{N}(\mathbf{z}_{k+1} - \boldsymbol{\mu}_{k+1|k}^i, \mathbf{H}_k^{(i)} \mathbf{P}_{k+1|k}^{(i)} (\mathbf{H}_k^{(i)})^T + \mathbf{R}_k)$ for EKF and $\gamma^i \sim \mathcal{N}(\mathbf{z}_{k+1} - \boldsymbol{\mu}_{k+1|k}^i, \hat{\mathbf{S}}_{k+1})$ for UKF.

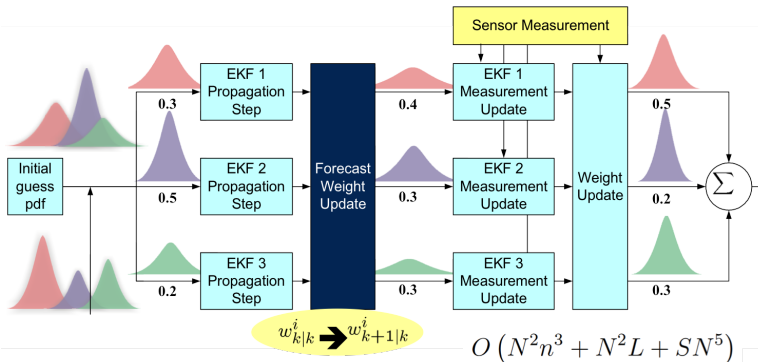
$$\mathbf{H}_k = \frac{\partial}{\partial \mathbf{x}_k} \mathbf{h}(t_k, \mathbf{x}_k) |_{\mathbf{x}_k = \boldsymbol{\mu}_{k+1|k}} \quad (5)$$

$\hat{\mathbf{S}}_{k+1}$ = innovation covariance¹

¹Julier, S., and Uhlmann, J., "Unscented Filtering and Nonlinear Estimation"

FILTERING UNDER BAYESIAN FRAMEWORK

ADAPTIVE GAUSSIAN SUM FILTER (AGSF)



- Highly Parallelized framework for Bayesian Nonlinear Filtering.

ADAPTIVE GAUSSIAN SUM FILTER

EXAMPLE:TWO BODY PROBLEM

Measurements are available when the satellite is in the field of view of the radar of an observation center. For simulation purposes, we consider the observation center to be located at 39.007° latitude and 104.883° longitude near Air Force Academy in Colorado springs. The cartesian coordinates of this location is given as:

$$\mathbf{r}_{\text{Site}} = \begin{bmatrix} -1275.1219 & -4797.9890 & 3994.2975 \end{bmatrix} \text{ km}$$

It is assumed that the measurements are available after 5.6585 hours i.e. time at which the last time-update was made. Six different cases are discussed

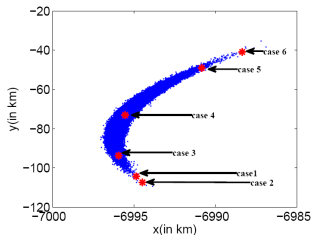
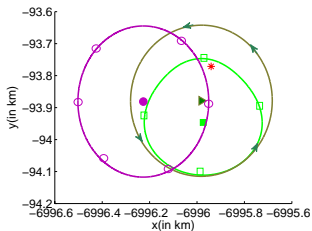


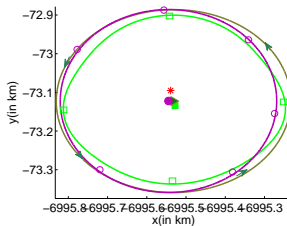
Figure: Different measurement cases. Case 4 corresponds to prior mean

ADAPTIVE GAUSSIAN SUM FILTER

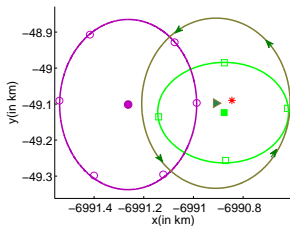
EXAMPLE: TWO BODY PROBLEM



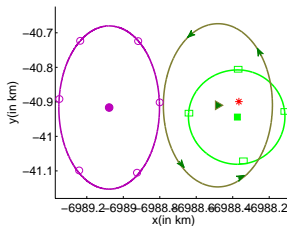
(a) case 3



(b) case 4 (measurements = prior mean)



(c) case 5



(d) case 6

ADAPTIVE GAUSSIAN SUM FILTER

EXAMPLE:TWO BODY PROBLEM

Filters	Pos. Error In Position (km)	Pos. Variance of Position (km^2)
AGSF	0.237	0.7837
UKF	0.244	0.0339
EKF	1.3329	0.03194

TABLE: Norm of error/variance when only position measurements are available and they are not at prior mean (Pos. Error = Posterior Error, Pos. Variance = Posterior Variance)²

The EKF provides completely inconsistent estimates of orbital states.

²K. Vishwajeet, P. Singla and M. Jah, "Nonlinear Uncertainty Propagation for Perturbed Two-Body Orbits," *AIAA Journal of Guidance, Control and Dynamics*, January 2014, DOI: 10.2514/1.G000472.

ADAPTIVE GAUSSIAN SUM FILTER

EXAMPLE: TWO BODY PROBLEM

Scenario II: Range, azimuthal and elevation angles are available for measurement

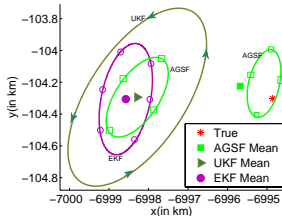
Measurement model is:

$$\mathbf{z}_{k+1}(1) = \sqrt{(\mathbf{x}_{k+1}(1) - \mathbf{rSite}(1))^2 + (\mathbf{x}_{k+1}(2) - \mathbf{rSite}(2))^2 + (\mathbf{x}_{k+1}(3) - \mathbf{rSite}(3))^2} + \mathbf{v}_1$$

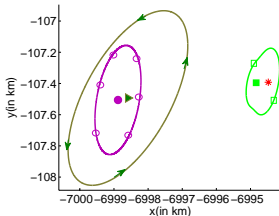
$$\mathbf{z}_{k+1}(2) = \tan^{-1} \frac{\mathbf{x}_{k+1}(3) - \mathbf{rSite}(3)}{\sqrt{(\mathbf{x}_{k+1}(1) - \mathbf{rSite}(1))^2 + (\mathbf{x}_{k+1}(2) - \mathbf{rSite}(2))^2}} + \mathbf{v}_2$$

$$\mathbf{z}_{k+1}(3) = \tan^{-1} \frac{\mathbf{x}_{k+1}(2) - \mathbf{rSite}(2)}{\mathbf{x}_{k+1}(1) - \mathbf{rSite}(1)} + \mathbf{v}_3$$

where, $\mathbf{v} = \begin{bmatrix} \mathbf{v}_1 & \mathbf{v}_2 & \mathbf{v}_3 \end{bmatrix}^T \sim \mathcal{N}(\mathbf{0}, \mathbf{R}_2)$ and, $\mathbf{R}_2 = \text{diag} \left(\underbrace{0.01}_{\text{km}^2} \quad \underbrace{0.0174 \quad 0.0174}_{\text{rad}^2} \right)$



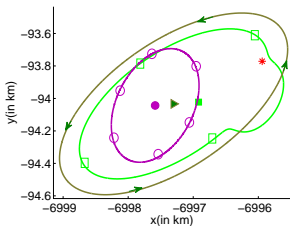
(a) case 1



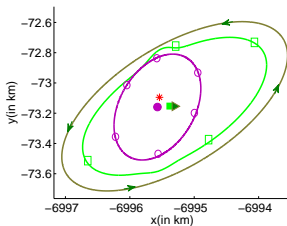
(b) case 2

ADAPTIVE GAUSSIAN SUM FILTER

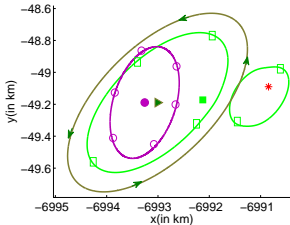
EXAMPLE: TWO BODY PROBLEM



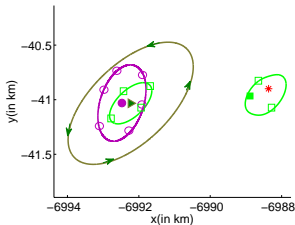
(a) case 3



(b) case 4 (measurements = prior mean)



(c) case 5



(d) case 6

ADAPTIVE GAUSSIAN SUM FILTER

EXAMPLE:TWO BODY PROBLEM

Filters	Pos. Error In Position (km)	Pos. Variance of Position (km^2)
AGSF	2.042	3.4608
UKF	7.1679	3.1315
EKF	7.7479	0.3987

TABLE: Norm of error when range and angular measurements are available and they are not at prior mean (Pos. Error = Posterior Error, Pos. Variance = Posterior Variance)³

The EKF provides completely inconsistent estimates of orbital states.

³K. Vishwajeet, P. Singla and M. Jah, "Nonlinear Uncertainty Propagation for Perturbed Two-Body Orbits," *AIAA Journal of Guidance, Control and Dynamics*, January 2014, DOI: 10.2514/1.G000472.

NUMERICAL EXPERIMENTS

ICELAND VOLCANO (EYJAFJALLAJÖKULL) ERUPTION



- The **BENT integral eruption column model** was used to produce eruption column parameters (mass loading, column height, grain size distribution) given *a specific atmospheric sounding and source conditions*.
 - BENT takes into consideration atmospheric (wind) conditions as given by atmospheric sounding data.
 - Plume rise height is given as *a function of volcanic source and environmental conditions*.
- The **PUFF Lagrangian model** was used to propagate ash parcels in *a given wind field (NCEP Reanalysis)*.
 - PUFF also accounts for dry deposition as well as deposition of ash on the ground.
- Polynomial chaos quadrature (PCQ) was used to select sample points and weights in the uncertain input space of *vent radius, vent velocity, mean particle size and particle size variance*.

- The **BENT integral eruption column model** was used to produce eruption column parameters (mass loading, column height, grain size distribution) given *a specific atmospheric sounding and source conditions*.
 - BENT takes into consideration atmospheric (wind) conditions as given by atmospheric sounding data.
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TABLE: Eruption source parameters based on observations of Eyjafjallajökull volcano and information from other similar eruptions.

Parameter	Value range	PDF	Comment
Vent radius, b_0 , m	65-150	Uniform	Measured from radar image of summit vents
Vent velocity, w_0 , m/s	Range: 45-124	Uniform	M. Ripepe, Geneva, Switzerland, 2010, presentation
Mean grain size, Md_ϕ	3.5-7	Uniform	Woods and Bursik (1991), Table 1, vulcanian and phreatoplinian. A. Hoskuldsson, Iceland meeting 2010, presentation
σ_ϕ	0.5 - 3	Uniform	Woods and Bursik (1991), Table 1, vulcanian and phreatoplinian.

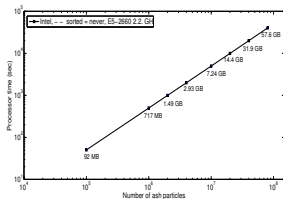
NUMERICAL EXPERIMENTS

ICELAND VOLCANO (EYJAFJALLAJÖKULL) ERUPTION

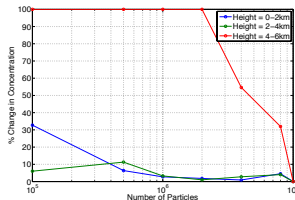
- 4×10^7 particles were used in BENT-PUFF model.

Location 52N, 13.5E: conc is the puff computed absolute air concentration (in mg/m^3) in a grid cell of size $0.5^\circ \times 0.5^\circ \times 2km$ at 1200hours on 16th April, 2010, and count is the number of PUFF particles in that cell

# of particles	height (km)	conc. $\times 10^{-4}$	count	height	conc. $\times 10^{-5}$	count	height	conc. $\times 10^{-7}$	count
10^5	3	0.74	28	5	4.23	16	7	-	-
5×10^5	3	1.17	221	5	3.54	67	7	-	-
10^6	3	1.12	405	5	4.12	156	7	-	-
2×10^6	3	1.12	884	5	4.03	305	7	-	-
4×10^6	3	1.09	1655	5	4.10	3620	7	1.32	2
8×10^6	3	1.15	3471	5	4.15	1256	7	1.98	6
10^7	3	1.10	4151	5	3.99	1510	7	2.91	11



Processor Time for Single Deterministic Run of BENT-PUFF vs. Number of Ash Particles.



Concentration (52N 13.5E) vs. Number of PUFF Particles

MEAN ASH TOP HEIGHT

ICELAND VOLCANO (EYJAFJALLAJÖKULL) ERUPTION

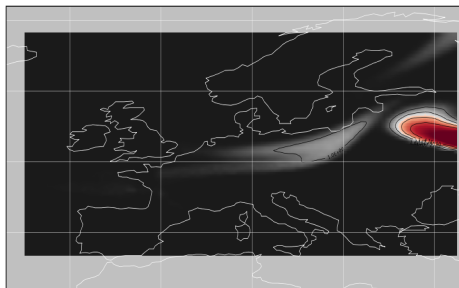


Figure: 9^4 Clenshaw Curtis Runs

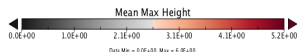
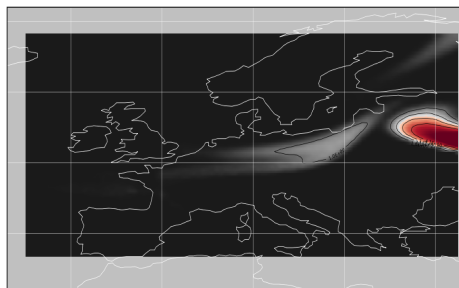


Figure: 161 CUT Runs

STANDARD DEVIATION OF ASH TOP HEIGHT

ICELAND VOLCANO (EYJAFJALLAJÖKULL) ERUPTION

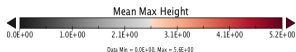
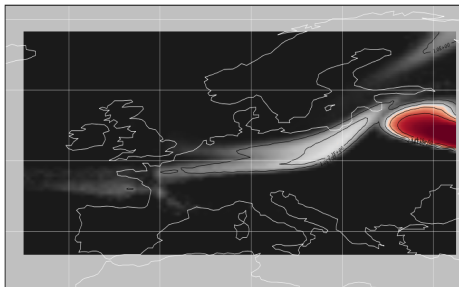


Figure: 9^4 Clenshaw Curtis Runs

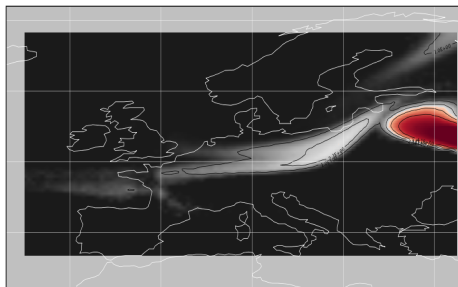
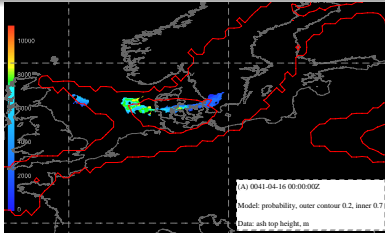


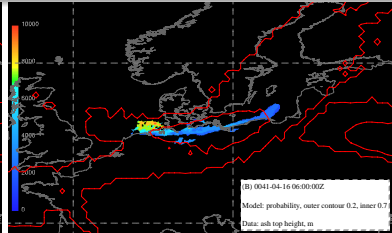
Figure: 161 CUT Runs

PROBABILITY OF ASH TOP HEIGHT

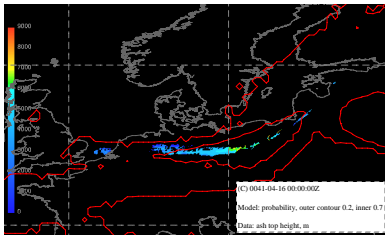
ICELAND VOLCANO (EYJAFJALLAJÖKULL) ERUPTION



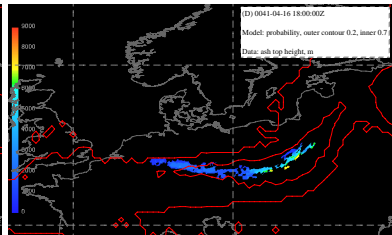
(a) 00 hrs



(b) 06 hrs



(c) 12 hrs



(d) 18 hrs

R. Madankan, et al., "Computation of Probabilistic Hazard Maps and Source Parameter Estimation For Volcanic Ash Transport

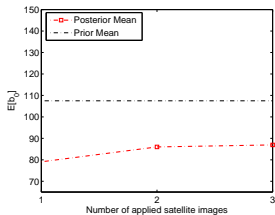
and Diffusion." *Journal of Computational Physics*. DOI: 10.1016/j.jcp.2013.11.032



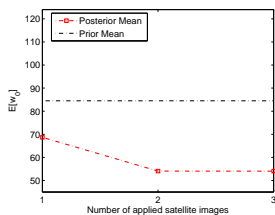
- Ash top-height (obtained from satellite imagery) is used as measurement data.
- Satellite data from three different time instants (April 16th at 0600 hrs, 1200 hrs, and 1800 hrs) are used as measurement data.
- Satellite observed ash top-heights are *assumed* to be accurate to within 100 m intervals around the observed height.
- Due to height quantization in the bent-puff model, ash top-height provided by bent-puff model is assumed to be polluted with zero-mean uniformly distributed random noise between -1000 m and $+1000\text{ m}$.
- Minimum variance framework was used for source parameter estimation.

APPLICATION TO ASH DISPERSION PROBLEM

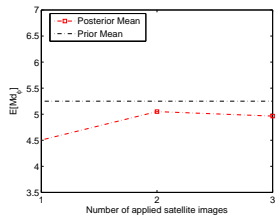
SOURCE PARAMETER ESTIMATION



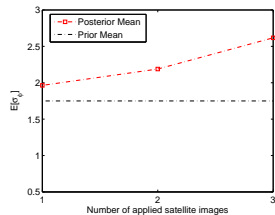
(e) Vent Radius



(f) Vent Velocity



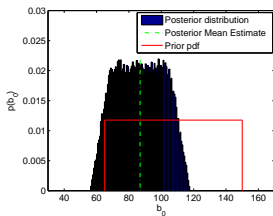
(g) Mean Grain Size



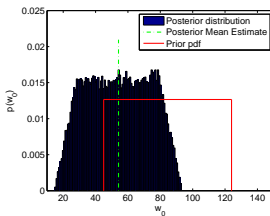
(h) σ_ϕ

APPLICATION TO ASH DISPERSION PROBLEM

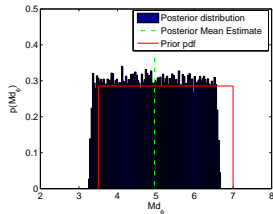
SOURCE PARAMETER ESTIMATION



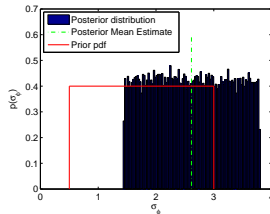
(i) Vent Radius



(j) Vent Velocity



(k) Mean Grain Size

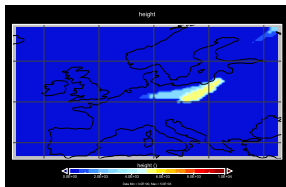


(l) σ_p

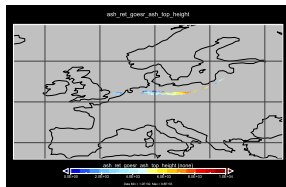
APPLICATION TO ASH DISPERSION PROBLEM

ASH PLUME FORECASTING

Comparison of Forecast of Ash top-height and Satellite Observation on April 16th, 1200 hrs.

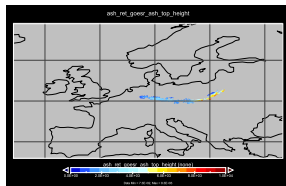
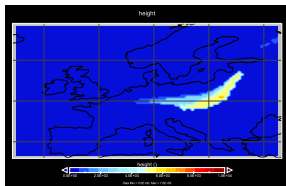


(a) Model Forecast



(b) Satellite Observation

Comparison of Forecast of Ash top-height and Satellite Observation on April 16th, 1800 hrs.



- **Non-Gaussian Approximation:**

- AGSF > Quadrature & CUT based filters > UKF > EKF

- **Computational Cost:**

- AGSF > Quadrature & CUT based filters > UKF ~ EKF

- **Ease of Implementation:**

- Quadrature & CUT based filters ~ UKF > EKF > AGSF

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